

## Unit 2 Unit-2

### Mahalanobis D. square

Mahalanobis distance is the distance between two points in multivariate space. In regular Euclidean space, variables ( $x, y, z$ ) are represented by axes drawn at right angles to each other. The distance between any two can be measured with a ruler. For uncorrelated variables, Euclidean distance equals the MD. However, if two or more variables are correlated, the axes are no longer at right angle and the measurement with becomes impossible with a ruler. In addition, if you have more than three variables, you can't plot them in regular 3D space at all. MD solves it.

this measurement problem, as it measures distances between points, even correlated points for multiple variables.

### Mahalanobis distance

The Mahalanobis distance measures distance relative to centroid - a base or central point which can be thought of as an overall mean for multivariate data. The centroid is a point in multivariate space where all means from all variables intersect. The larger the MD, the further away from the centroid the data point is.

### Application

MD is widely used in cluster analysis and classification techniques. It is closely related to T-square distance distribution used for multivariate statistical testing and Fisher's linear discriminant analysis that is used for supervised classification.

MD is

In order to use the Mahalanobis distance to classify a test point as belonging to one of  $N$  classes, one first estimates the covariance matrix of each class, usually based on samples known to belong to each class. Then, given a test sample, one computes the Mahalanobis distance to each class and classifies the test point as belonging to that class for which the Mahalanobis distance is minimal.

Mahalanobis distance and leverage are often used to detect outliers especially in the development of linear regression models. A point that has a greater Mahalanobis distance from the rest of the sample population of points is said to have higher leverage since it has a higher greater influence on the slope or coefficients of the regression equation.

MD is also used to determine multivariate outliers. Regression techniques can be used to determine if a specific case within a sample population is an outlier via the combination of two or more variables scores. Even for normal distributions, a point can be a multivariate outlier even if it is not a univariate outlier for any variables.

### Formal definition

The MD between two objects is defined as

$$D^2(\text{Mahalanobis}) = [(x_B - x_A)^T * C^{-1} * (x_B - x_A)]$$

where  $x_A$  and  $x_B$  is a pair of objects

$C$  is the sample covariance matrix.

Another version of the formula, which uses distances from each observation to the central mean

$$d_i = [(x_i - \bar{x})^T * C^{-1} * (x_i - \bar{x})]^{0.5}$$



where

$x$  = an object vector

$\bar{x}$  = A.M vector

$C$  = sample covariance matrix.

### Disadvantages

Although MD is included with many popular stat packages, some authors question the reliability of results.

A major issue with the MD is that the inverse of the correlation matrix is needed for the calculations. This can't be calculated if the variables are highly correlated.

### Generalised $T^2$ statistic distribution.

For univariate distribution,

Recalling the univariate theory for determining whether a specific value  $\mu_0$  is a plausible value for the population mean  $\mu$ .

Null hypothesis:  $\mu = \mu_0$  Alternative hypothesis  $\mu \neq \mu_0$

If  $x_1, x_2, \dots, x_n$  denote random sample from normal population, then the test statistic is

$$t = \frac{(\bar{x} - \mu_0)}{s/\sqrt{n}} \quad \text{where } \bar{x} = \frac{\sum x_i}{n} \quad s^2 = \frac{1}{n-1} \sum (x_i - \bar{x})^2$$

$t \sim t$  distribution with  $n-1$  d.f.

If  $|t| > t_{n-1}(\alpha/2)$ , reject  $H_0$ .

Rejecting  $H_0$  when  $|t|$  is large is equivalent to rejecting  $H_0$

if its square  $t^2 = \frac{(\bar{x} - \mu_0)^2}{s^2/n} = n(\bar{x} - \mu_0)(s^2)^{-1}(\bar{x} - \mu_0)$  — (1)

is large. The variable  $t^2$  is a square of the distance from the sample mean  $\bar{x}$  to test value  $\mu_0$ . The units of distance are expressed in terms of  $s/\sqrt{n}$ .

Reject  $H_0$  at significance level  $\alpha$ , if

$$n(\bar{x} - \mu_0)(s^2)^{-1}(\bar{x} - \mu_0) > t_{n-1}(\alpha/2) \quad \text{--- (2)}$$

where  $t_{n-1}(\alpha/2)$  denotes the upper  $100(\alpha/2)$ th percentile of the  $t$ -distribution with  $n-1$  d.f.

$$\{ \text{Accept } H_0 : \mu = \mu_0 \text{ at level } \alpha \} \text{ (or) } \left| \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \right| \leq t_{n-1}(\alpha/2)$$

is equivalent to

$\left\{ \mu_0 \text{ lies in the } 100(1-\alpha)\% \text{ confidence interval } \bar{x} \pm t_{n-1}(\alpha/2) \frac{s}{\sqrt{n}} \right\}$

(or)

$$\bar{x} - t_{n-1}(\alpha/2) \frac{s}{\sqrt{n}} \leq \mu_0 \leq \bar{x} + t_{n-1}(\alpha/2) \frac{s}{\sqrt{n}} \quad \text{--- (3)}$$

Confidence interval in (3) is a random interval because the end points depends upon the random variable  $\bar{x}$  and  $s$ .

A natural generalization of the squared distance in ① is its multivariate analog

$$T^2 = (\bar{x} - \mu_0)' \left( \frac{S}{n} \right)^{-1} (\bar{x} - \mu_0) = n (\bar{x} - \mu_0)' S^{-1} (\bar{x} - \mu_0) \quad \text{--- ④}$$

$$\begin{matrix} p \times 1 \\ \bar{x} = \frac{\sum_{j=1}^n x_j}{n} \end{matrix} \quad \begin{matrix} p \times p \\ S = \frac{1}{n-1} \sum_{j=1}^n (x_j - \bar{x})(x_j - \bar{x})' \end{matrix} \quad \begin{matrix} p \times 1 \\ \mu_0 = \begin{bmatrix} \mu_{10} \\ \mu_{20} \\ \vdots \\ \mu_{p0} \end{bmatrix} \end{matrix}$$

The statistic  $T^2$  is called Hotelling  $T^2$ .

$(1/n)S$  is the estimated covariance matrix of  $\bar{x}$ .

If  $T^2$  is too large,  $H_0$  is rejected.

$T^2$  is distributed as  $\frac{(n-1)p}{n-p} F_{p, n-p}$

$$\bar{x} = \frac{\sum x_j}{n} \quad S = \frac{1}{n-1} \sum_j (x_j - \bar{x})(x_j - \bar{x})'$$

$$\alpha = P \left[ T^2 > \frac{(n-1)p}{n-p} F_{p, n-p}(\alpha) \right]$$

$$\alpha = P \left[ n(\bar{x} - \mu)' S^{-1} (\bar{x} - \mu) > \frac{(n-1)p}{n-p} F_{p, n-p}(\alpha) \right]$$

$$\text{Reject } H_0 \text{ if } n(\bar{x} - \mu)' S^{-1} (\bar{x} - \mu) > \frac{(n-1)p}{n-p} F_{p, n-p}(\alpha)$$



$$T^2 = \sqrt{n} (\bar{x} - \mu_0)' \left( \frac{\sum_{j=1}^n (x_j - \bar{x})(x_j - \bar{x})'}{n-1} \right) \sqrt{n} (\bar{x} - \mu_0)$$

which combines normal  $N(0, \Sigma)$  random vector and Wishart  $W_{p, n-1}(\Sigma)$  random matrix in the form

$$T^2_{p, n-1} = \left( \begin{array}{c} \text{multivariate normal} \\ \text{random vector} \end{array} \right) \left( \begin{array}{c} \text{Wishart random} \\ \text{matrix} \\ \hline d.f. \end{array} \right) \left( \begin{array}{c} \text{multivariate normal} \\ \text{random vector} \end{array} \right)$$

This is analogous to

$$t^2 = \sqrt{n} (\bar{x} - \mu_0)' (s^2)^{-1} \sqrt{n} (\bar{x} - \mu_0)$$

$$t^2_{n-1} = \left( \begin{array}{c} \text{normal} \\ \text{random vector} \end{array} \right) \left( \begin{array}{c} \text{chi-squared} \\ \hline d.f. \end{array} \right)^{-1} \left( \begin{array}{c} \text{normal} \\ \text{random variable} \end{array} \right)$$

for univariate case

## Unit 2:- Sampling distribution of Covariance matrix

The sample covariance matrix  $s$ , 
$$s = \frac{1}{N-1} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})'$$

is an unbiased estimator of the popln covariance matrix  $\Sigma$ . This result will be generalised to the case of a matrix  $A$  of any order when  $\Sigma = I$ , this distribution is in a sense a generalisation of the  $\chi^2$ -distribution. The distribution of  $A$  (or  $s$ ), often called the Wishart distribution, is fundamental to multivariate statistical analysis.

The tentative assumption that  $x_1, x_2, \dots, x_n$  constitute a random sample from a normal popln with mean  $\mu$  and covariance  $\Sigma$  completely determines the sampling distributions of  $\bar{x}$  and  $s$ . Here we present the results on the sampling distn of  $\bar{x}$  and  $s$  by drawing a parallel with the familiar univariate conclusions.

In the univariate case ( $p=1$ ), W.K.T  $\bar{x}$  is normal with mean  $\mu =$  popln mean and variance

$$\frac{1}{n} \sigma^2 = \frac{\text{Population variance}}{\text{Sample size}}$$

The result for the multivariate case ( $p \geq 2$ ) is analogous in that  $\bar{x}$  has a normal distn with mean  $\mu$  and covariance matrix  $(\frac{1}{n}) \Sigma$ .



For the sample variance, recall that ~~(n-1)~~

$$(n-1) s^2 = \sum_{j=1}^n (x_j - \bar{x})^2 \text{ is distributed as } \sigma^2 \text{ times a}$$

chi-square variable having  $(n-1)$  degrees of freedom. In turn, this chi-square is the distribution of a sum of squares of independent standard normal random variables.

That is  $(n-1) s^2$  is distributed as

$$\sigma^2 (\chi_1^2 + \dots + \chi_{n-1}^2) = (\sigma \chi_1)^2 + (\sigma \chi_2)^2 + \dots + (\sigma \chi_{n-1})^2$$

The individual term  $\sigma \chi_i$  are independently distributed as  $N(0, \sigma^2)$

Wishart distribution

The sampling distribution of the sample covariance matrix is called the wishart distribution, after its discoverer, it is defined as the sum of independent products of multivariate normal random vectors. specifically,

$W_m(\cdot | \Sigma) =$  wishart distribution with  $m$  d.f

$$= \text{distribution of } \sum_{j=1}^m \tilde{x}_j \tilde{x}_j'$$

where  $\tilde{x}_j$  are each independently distributed as  $N_p(0, \Sigma)$

We summarise the sampling distn results as follows:-

Let  $x_1, x_2, \dots, x_n$  be a random sample of size  $n$  from a  $p$ -variate normal distn with mean  $\mu$  and covariance matrix  $\Sigma$ .

Then

1)  $\bar{x}$  is distributed  $N_p(\mu, (1/n)\Sigma)$

2)  $(n-1)s$  is distributed as a Wishart random matrix with  $(n-1)$  df

3)  $\bar{x}$  and  $s$  are independent

Because  $\Sigma$  is unknown, the distr of  $\bar{x}$  cannot be used directly to make inferences about  $\mu$ . However,  $s$  provides independent information about  $\Sigma$ , and the distr of  $s$  does not depend on  $\mu$ . This allows us to construct a statistic for making inferences about  $\mu$ .

Properties of the Wishart distribution

1) If  $A_1$  is distributed as  $W_{m_1}(A_1/\Sigma)$  independent of  $A_2$ , which is distributed as  $W_{m_2}(A_2/\Sigma)$ , then  $A_1 + A_2$  is distributed as  $W_{m_1+m_2}(A_1+A_2/\Sigma)$ . That is, the degrees of freedom add.

2) If  $A$  is distributed as  $W_m(A/\Sigma)$ , then  $CA C'$  is distributed as  $W_m(CA C' / C \Sigma C')$

The probability density function of the Wishart distribution, the sample size  $n$  is greater than the no. of variables  $p$ , then



$$|A|^{(n-p-2)/2} \frac{e^{-\text{tr}[A\Sigma^{-1}]/2}}{e}$$

$$W_{n-1}(A/\Sigma) = \frac{2^{p(n-1)/2} \pi^{p(p-1)/4} |\Sigma|^{(n-1)/2} \prod_{i=1}^p \Gamma\left(\frac{1}{2}(n-i)\right)}{e}$$

A positive definite where  $\Gamma(\cdot)$  is the gamma function.

## Uses and Application of $T^2$

- 1) Hotelling T-square statistic allows for the testing of hypotheses on multiple (often correlated) measures within the same sample.
- 2) A generalization of student's t-statistic, called Hotelling  $T^2$  statistic, allows for the testing of hypotheses on multiple (often correlated) measures within the same sample.
- 3) Hotelling's T-squared has several advantages over the t-test
- 4) The type I error rate is well controlled,
- 5) The relationship b/w multiple variables is taken into account.
- 6) It can generate an overall conclusion even if multiple (single) t-tests are inconsistent. While a t-test will test you which variable differ b/w groups, Hotelling's summarizes the b/w-group differences.
- 7) From the Hotelling's T-squared test it was concluded that there is significant difference in the average performance of boarding and day students. On the other hand, the 2-way ANOVA indicates a significant difference in the performance of students based on subjects and then gives no evidence to conclude that there is significant difference b/w the years under study



Case I: Inference problems regarding  $\mu$  means from a paired popn

Case II: Inference problems regarding a single multivariate popn

Case III: Inference problems regarding  $\mu$  means from 2 independent popn

All three cases make use of the Hotelling T-square statistic

8) In statistics, particularly in hypothesis testing, the Hotelling T-squared distribution ( $T^2$ ), proposed by Harold Hotelling, is a multivariate prob distr that is tightly related to the F-distr and is most notable for arising as the distr of a set of sample statistics that are natural generalizations of the statistics underlying the student's t-distr.

9) The Hotelling's T-square test plays the same role as the t-test when there is more than one dependent variable. It is a special case of MANOVA but not of ANOVA.

One sample hypothesis testing of the mean

• sample hypothesis testing of the mean with paired samples

• sample hypothesis testing of the mean with independent samples

• sample hypothesis testing of means with unequal covariance matrices

~~2-sample hypothesis~~

Real statistics capabilities

statistical power and sample size requirements

Hotelling's T-squared is based on Hotelling's  $T^2$  distr and forms the

basis for various multivariate control charts